

PROBABILITY DENSITY FUNCTIONS OF ORDER STATISTICS OF NONIDENTICALLY DISTRIBUTED TRUNCATED CONTINUOUS RANDOM VARIABLES**Gökhan Gökdere**

Bingol University, Department of Mathematics, Bingol, Turkey

Abstract

In this paper, the theorems related to the probability density functions of order statistics of independent but not necessarily identically distributed truncated continuous random variables are given. Then some results are obtained.

Keywords: Order statistics, joint distribution, truncated distribution, permanent.

1. Introduction

A good deal of work has appeared in the literature about several identities and recurrence relations for probability density function (pdf) and distribution function (df) of order statistics of independent but not necessarily identically distributed (innid) random variables. [1] derived the joint pdf and marginal pdf of order statistics of innid random variables by means of permanents. [2] obtained the distribution of single order statistics in terms of distribution functions of the minimum and maximum order statistics of some subsets of $\{X_1, X_2, \dots, X_n\}$ where X_i 's are innid random variables. Several recurrence relations and identities for product moments of order statistics of innid random variables by using permanents were obtained [3]. Some expressions for the moments and product moments of the order statistics from the doubly truncated linear-exponential distributions were derived [4]. [5] obtained the recurrence relationships among the distribution function of order statistics arising from innid random variables. [6] generalized the independent and identically distributed (iid) results for the pareto and doubly-truncated pareto models established by [7]. [8] derived some general recurrence relations satisfied by single and product moments of order statistics from doubly truncated continuous distributions. [9] obtained the joint pdf and df of order statistics of innid random variables by means of permanents.

Let F_1, F_2, \dots, F_n and f_1, f_2, \dots, f_n be df and pdf of X_1, X_2, \dots, X_n , respectively. In addition, the inverse function of F_1, F_2, \dots, F_n are expressed as

$$F_i^{-1}(p) = \inf\{x : F_i(x) \geq p\} = \sup\{x : F_i(x) < p\}, \quad (1)$$

where $x \in R$, $0 < p < 1$ and $i = 1, 2, \dots, n$ [10,11]. In (1), if $p = 0$ and then $p = 1$, the following equations are obtained [12].

$$\alpha(F_i) = \inf\{x : F_i(x) > 0\} = F_i^{-1}(0) \quad \text{and} \quad w(F_i) = \sup\{x : F_i(x) < 1\} = F_i^{-1}(1). \quad (2)$$

From (2), we can write

$$\alpha({}_{uv}F_i) = u_i \quad \text{ve} \quad w({}_{uv}F_i) = v_i. \quad (3)$$

From (3), df and pdf of X_i which truncated on the left at u_i and right at v_i , respectively, are expressed as

$${}_{uv}F_i(x) = \frac{F_i(x) - F_i(u_i)}{F_i(v_i) - F_i(u_i)} \quad \text{and} \quad {}_{uv}f_i(x) = \frac{f_i(x)}{F_i(v_i) - F_i(u_i)} \quad (4)$$

Let X_1, X_2, \dots, X_n be innid continuous random variables which are truncated and let $X_{1:n} \leq X_{2:n} \leq \dots \leq X_{n:n}$ denote the corresponding order statistics.

In addition, $X^s_{1:n}, X^s_{2:n}, \dots, X^s_{n:n}$ are order statistics of truncated random variables with common df ${}_{uv}F^s$ and pdf ${}_{uv}f^s$, respectively, defined by

$${}_{uv}F^s = \frac{1}{n_s} \sum_{i \in s} {}_{uv}F_i \quad (5)$$

and

$${}_{uv}f^s = \frac{1}{n_s} \sum_{i \in s} {}_{uv}f_i \quad (6)$$

where s is a subset of integers $\{1, 2, \dots, n\}$ with $n_s \geq 1$ elements. Moreover, if a_1, a_1, \dots are column vectors,

$$A = [a_1, a_2, \dots]$$

then $\begin{matrix} i_1 & i_2 & \dots \end{matrix}$ will denote the matrix obtained by taking i_1 copies of a_1 , i_2 copies of a_2 and so on. Finally, $\text{Per}(A)$ denotes the permanent of a square matrix A , which is defined similar to the determinants except that all terms in the expansion have a positive sign [13].

2. Main Theorems

In this section, we obtain theorems related to the probability density functions of order statistics of innid truncated continuous random variables. Now the pdf of $X_{r_1:n}, X_{r_2:n}, \dots, X_{r_d:n}$ ($1 \leq r_1 < r_2 < \dots < r_d \leq n$, $d = 1, 2, \dots, n$) will be given.

Theorem 2.1.

$$f_{r_1, r_2, \dots, r_d; n}(x_1, x_2, \dots, x_d) = D \sum_{t_d, \dots, t_2, t_1}^{n, \dots, r_3-1, r_2-1} (-1)^{-d + \sum_{w=1}^d (r_{w+1} - t_w)} \prod_{w=1}^d \binom{r_{w+1} - r_w - 1}{t_w - r_w} \sum_{n_s = n + r_d - t_d} (t_d - r_d)! \sum_{n_{s_1}, n_{s_2}, \dots, n_{s_{d-1}}} \prod_{w=1}^d \sum_{n_{s_w}} \text{per} [\begin{matrix} {}_{uv}F(x_w) \\ r_{w+1} - r_{w-1} - 1 - t_w + t_{w-1} \end{matrix}] [\zeta_w / \cdot] \text{per} [\begin{matrix} {}_{uv}f(x_w) \\ 1 \end{matrix}] [\zeta_w' / \cdot] \quad (7)$$

where $D = \left[\prod_{w=1}^{d+1} (r_w - r_{w-1} - 1)! \right]^{-1}$, $r_0 = 0$, $r_{d+1} = n + 1$, $\sum_{t_d, \dots, t_2, t_1}^{n, \dots, r_3-1, r_2-1} = \sum_{t_d=r_d}^n \dots \sum_{t_2=r_2}^{r_3-1} \sum_{t_1=r_1}^{r_2-1}$, $s = \bigcup_{w=1}^d s_w$, $v \neq v'$, $s_v \cap s_{v'} = \emptyset$, $s = s_w \cup s_w'$, $n_{s_w} = r_{w+1} - r_{w-1} - t_w + t_{w-1}$, $t_0 = r_1 - 1$, $n_{s_w} = r_{w+1} - r_{w-1} - 1 - t_w + t_{w-1}$ and $n_{s_w'} = 1$. In addition, ${}_{uv}f(x_w) = ({}_{uv}f_1(x_w), {}_{uv}f_2(x_w), \dots, {}_{uv}f_n(x_w))'$ is column vector.

Proof. Let us consider

$$P\{x_1 < X_{r_1:n} \leq x_1 + \delta x_1, x_2 < X_{r_2:n} \leq x_2 + \delta x_2, \dots, x_d < X_{r_d:n} \leq x_d + \delta x_d\} \tag{8}$$

Dividing (8) to $\prod_{w=1}^d \delta x_w$ and $\delta x_1 \rightarrow 0, \delta x_2 \rightarrow 0, \dots, \delta x_d \rightarrow 0$ we obtain

$${}_{uv}f_{r_1, r_2, \dots, r_d:n}(x_1, x_2, \dots, x_d) = D \text{ per} B \tag{9}$$

where $B = \begin{bmatrix} {}_{uv}F(x_1) & {}_{uv}f(x_1) & {}_{uv}F(x_2) & -{}_{uv}F(x_1) & {}_{uv}f(x_2) & \dots & {}_{uv}f(x_d) & 1 - {}_{uv}F(x_d) \\ r_1-1 & 1 & r_2-r_1-1 & 1 & 1 & \dots & 1 & n-r_d \end{bmatrix}$ is a square matrix.

Using properties of permanent, it can be written

$$\begin{aligned} \text{per} B &= \text{per} \begin{bmatrix} {}_{uv}F(x_1) & {}_{uv}f(x_1) & {}_{uv}F(x_2) & -{}_{uv}F(x_1) & {}_{uv}f(x_2) & \dots & {}_{uv}f(x_d) & 1 - {}_{uv}F(x_d) \\ r_1-1 & 1 & r_2-r_1-1 & 1 & 1 & \dots & 1 & n-r_d \end{bmatrix} \\ &= \sum_{t_d=0}^{n-r_d} (-1)^{n-r_d-t_d} \binom{n-r_d}{t_d} \dots \sum_{t_2=0}^{r_3-r_2-1} (-1)^{r_3-r_2-1-t_2} \binom{r_3-r_2-1}{t_2} \sum_{t_1=0}^{r_2-r_1-1} (-1)^{r_2-r_1-1-t_1} \binom{r_2-r_1-1}{t_1} \\ &\quad \cdot \text{per} \begin{bmatrix} {}_{uv}F(x_1) & {}_{uv}f(x_1) & {}_{uv}F(x_2) & {}_{uv}f(x_2) & \dots & {}_{uv}f(x_d) & 1 & {}_{uv}F(x_d) \\ r_2-2-t_1 & 1 & r_3-r_2-1-t_2+t_1 & 1 & \dots & 1 & t_d & n-r_d-t_d+t_{d-1} \end{bmatrix} \\ &= \sum_{t_d=0}^{n-r_d} \dots \sum_{t_2=0}^{r_3-r_2-1} \sum_{t_1=0}^{r_2-r_1-1} (-1)^{n+1-r_1-d-\sum_{w=1}^d t_w} \prod_{w=1}^d \binom{r_{w+1}-r_w-1}{t_w} \sum_{n_s=n-t_d} t_d! \\ &\quad \cdot \text{per} \begin{bmatrix} {}_{uv}F(x_1) & {}_{uv}F(x_2) & \dots & {}_{uv}F(x_d) & {}_{uv}f(x_1) & {}_{uv}f(x_2) & \dots & {}_{uv}f(x_d) \\ r_2-2-t_1 & r_3-r_2-1-t_2+t_1 & \dots & n-r_d-t_d+t_{d-1} & 1 & 1 & \dots & 1 \end{bmatrix} [s/.] \\ &= \sum_{t_d=r_d}^n \dots \sum_{t_2=r_2}^{r_3-1} \sum_{t_1=r_1}^{r_2-1} (-1)^{-d+\sum_{w=1}^d (r_{w+1}-t_w)} \prod_{w=1}^d \binom{r_{w+1}-r_w-1}{t_w-r_w} \sum_{n_s=n+r_d-t_d} (t_d-r_d)! \sum_{n_{s_1}, n_{s_2}, \dots, n_{s_{d-1}}} \\ &\quad \cdot \text{per} \begin{bmatrix} {}_{uv}F(x_1) & {}_{uv}F(x_2) & \dots & {}_{uv}F(x_d) & {}_{uv}f(x_1) & {}_{uv}f(x_2) & \dots & {}_{uv}f(x_d) \\ r_2+r_1-2-t_1 & r_3-r_1-1-t_2+t_1 & \dots & n-r_{d-1}-t_d+t_{d-1} & 1 & 1 & \dots & 1 \end{bmatrix} [s_1/.] \\ &= \sum_{t_d=r_d}^n \dots \sum_{t_2=r_2}^{r_3-1} \sum_{t_1=r_1}^{r_2-1} (-1)^{-d+\sum_{w=1}^d (r_{w+1}-t_w)} \prod_{w=1}^d \binom{r_{w+1}-r_w-1}{t_w-r_w} \sum_{n_s=n+r_d-t_d} (t_d-r_d)! \sum_{n_{s_1}, n_{s_2}, \dots, n_{s_{d-1}}} \\ &\quad \cdot \text{per} \begin{bmatrix} {}_{uv}F(x_1) & {}_{uv}f(x_1) \\ r_2+r_1-2-t_1 & 1 \end{bmatrix} [s_1/.] \text{per} \begin{bmatrix} {}_{uv}F(x_2) & {}_{uv}f(x_2) \\ r_3-r_1-1-t_2+t_1 & 1 \end{bmatrix} [s_2/.] \dots \text{per} \begin{bmatrix} {}_{uv}F(x_d) & {}_{uv}f(x_d) \\ n-r_{d-1}-t_d+t_{d-1} & 1 \end{bmatrix} [s_d/.] \\ &= \sum_{t_d=r_d}^n \dots \sum_{t_2=r_2}^{r_3-1} \sum_{t_1=r_1}^{r_2-1} (-1)^{-d+\sum_{w=1}^d (r_{w+1}-t_w)} \prod_{w=1}^d \binom{r_{w+1}-r_w-1}{t_w-r_w} \sum_{n_s=n+r_d-t_d} (t_d-r_d)! \sum_{n_{s_1}, n_{s_2}, \dots, n_{s_{d-1}}} \\ &\quad \cdot \prod_{w=1}^d \text{per} \begin{bmatrix} {}_{uv}F(x_w) & {}_{uv}f(x_w) \\ r_{w+1}-r_{w-1}-1-t_w+t_{w-1} & 1 \end{bmatrix} [s_w/.] \\ &= \sum_{t_d=r_d}^n \dots \sum_{t_2=r_2}^{r_3-1} \sum_{t_1=r_1}^{r_2-1} (-1)^{-d+\sum_{w=1}^d (r_{w+1}-t_w)} \prod_{w=1}^d \binom{r_{w+1}-r_w-1}{t_w-r_w} \sum_{n_s=n+r_d-t_d} (t_d-r_d)! \sum_{n_{s_1}, n_{s_2}, \dots, n_{s_{d-1}}} \end{aligned}$$

$$\prod_{w=1}^d \sum_{n_{\zeta_w}} \text{per} \left[\begin{matrix} uvF(x_w) \\ r_{w+1}-r_{w-1}-1-t_w+t_{w-1} \end{matrix} \right] [\zeta_w / \cdot] \text{per} \left[\begin{matrix} uvf(x_w) \\ 1 \end{matrix} \right] [\zeta'_w / \cdot] \tag{10}$$

After substituting (10) into (9), we can obtained (7).

Theorem 2.1. can be expressed as Theorem 2.2. using properties of permanent.

Theorem 2.2.

$$uvf_{r_1, r_2, \dots, r_d; n}(x_1, x_2, \dots, x_d) = D \sum_{t_d, \dots, t_2, t_1}^{n, \dots, r_3-1, r_2-1} (-1)^{-d + \sum_{w=1}^d (r_{w+1} - t_w)} \prod_{w=1}^d \binom{r_{w+1} - r_w - 1}{t_w - r_w} \sum_{n_s = n + r_d - t_d} (t_d - r_d)! \sum_{n_1, n_2, \dots, n_{s-d-1}} \prod_{w=1}^d \sum_{n_{\zeta_w}} n_{\zeta_w}! \prod_{l=1}^{n_{\zeta_w}} uvF_{\zeta_w^l}(x_w) uvf_{\zeta_w^{t_w}}(x_w) \tag{11}$$

where $\zeta_w = \{\zeta_w^1, \zeta_w^2, \dots, \zeta_w^{n_{\zeta_w}}\}$ and $\zeta'_w = \{\zeta_w^{t_w}\}$.

Theorem 2.2. can be expressed as Theorem 2.3. using (5) and (6).

Theorem 2.3.

$$uvf_{r_1, r_2, \dots, r_d; n}(x_1, x_2, \dots, x_d) = \sum \sum n! D \sum_{t_d, \dots, t_2, t_1}^{n, \dots, r_3-1, r_2-1} (-1)^{-d + \sum_{w=1}^d (r_{w+1} - t_w)} \prod_{w=1}^d \binom{r_{w+1} - r_w - 1}{t_w - r_w} [uvF^S(x_w)]^{r_{w+1} - r_{w-1} - 1 - t_w + t_{w-1}} uvf^S(x_w) \tag{12}$$

$$\sum \sum = \sum_{\kappa=1}^n (-1)^{n-\kappa} \frac{\kappa^n}{n!} \sum_{n_s = \kappa}$$

where

Proof. (8) can be expressed as

$$\sum \sum P\{x_1 < X_{r_1; n}^s \leq x_1 + \delta x_1, x_2 < X_{r_2; n}^s \leq x_2 + \delta x_2, \dots, x_d < X_{r_d; n}^s \leq x_d + \delta x_d\} \tag{13}$$

Dividing (13) to $\prod_{w=1}^d \delta x_w$ and $\delta x_1 \rightarrow 0, \delta x_2 \rightarrow 0, \dots, \delta x_d \rightarrow 0$ we obtain (12).

Theorem 2.1. can be expressed as Theorem 2.4. using expansion of permanent.

Theorem 2.4.

$$uvf_{r_1, r_2, \dots, r_d; n}(x_1, x_2, \dots, x_d) = D \sum_P \prod_{l=1}^{r_1-1} uvF_{j_l}(x_1) \prod_{w=2}^{d+1} \sum_{t=r_{w-1}}^{r_w-1} (-1)^{r_w-1-t} \sum_{\substack{n_\tau = t - r_{w-1} \\ n_{\tau'} = r_w - 1 - t}} \prod_{l=1}^{t-r_{w-1}} uvF_{\tau_l}(x_w) \prod_{l=1}^{r_w-1-t} uvF_{\tau'_l}(x_{w-1}) \prod_{w=1}^d uvf_{j_{r_w}}(x_w) \tag{14}$$

where \sum_P denotes the sum over all $n!$ permutations $\{j_1, j_2, \dots, j_n\}$ of $\{1, 2, \dots, n\}$, $\tau = \{\tau_1, \tau_2, \dots, \tau_{t-r_{w-1}}\}$, $\tau' = \{\tau'_1, \tau'_2, \dots, \tau'_{r_w-1-t}\}$, $\tau \cap \tau' = \emptyset$ and $\tau \cup \tau' = \{j_{r_{w-1}+1}, j_{r_{w-1}+2}, \dots, j_{r_w-1}\}$.

Proof. Let us consider

$$P\{x_1 < X_{r_1:n} \leq x_1 + \delta x_1, x_2 < X_{r_2:n} \leq x_2 + \delta x_2, \dots, x_d < X_{r_d:n} \leq x_d + \delta x_d\} \tag{15}$$

Dividing (15) to $\prod_{w=1}^d \delta x_w$ and $\delta x_1 \rightarrow 0, \delta x_2 \rightarrow 0, \dots, \delta x_d \rightarrow 0$ we obtain

$$uvf_{r_1, r_2, \dots, r_d:n}(x_1, x_2, \dots, x_d) = D \sum_P \prod_{l=1}^{r_1-1} [uvF_{j_l}(x_1)] uvf_{j_{r_1}}(x_1) \prod_{l=r_1+1}^{r_2-1} [uvF_{j_l}(x_2) - uvF_{j_l}(x_1)] uvf_{j_{r_2}}(x_2) \dots uvf_{j_{r_d}}(x_d) \prod_{l=r_d+1}^n [1 - uvF_{j_l}(x_d)] \tag{16}$$

(16) can be written,

$$uvf_{r_1, r_2, \dots, r_d:n}(x_1, x_2, \dots, x_d) = D \sum_P \prod_{l=1}^{r_1-1} uvF_{j_l}(x_1) \prod_{w=2}^{d+1} \prod_{l=r_{w-1}+1}^{r_w-1} [uvF_{j_l}(x_w) - uvF_{j_l}(x_{w-1})] \prod_{w=1}^d uvf_{j_{r_w}}(x_w) \tag{17}$$

Now, let us consider the following equation

$$\prod_{l=r_{w-1}+1}^{r_w-1} [uvF_{j_l}(x_w) - uvF_{j_l}(x_{w-1})] = \sum_{t=r_{w-1}}^{r_w-1} (-1)^{r_w-1-t} \sum_{\substack{n_\tau=t-r_{w-1} \\ n_{\tau'}=r_w-1-t}} \prod_{l=1}^{t-r_{w-1}} uvF_{\tau_l}(x_w) \prod_{l=1}^{r_w-1-t} uvF_{\tau'_l}(x_{w-1}) \tag{18}$$

After substituting (18) into (17), we can obtained (14). Theorem 2.4. can be expressed as Theorem 2.5.

Theorem 2.5.

$$uvf_{r_1, r_2, \dots, r_d:n}(x_1, x_2, \dots, x_d) = D^{-1} \sum_{D \in P_{r_d, \dots, r_2, n}} \prod_{l=1}^{r_1-1} uvF_{j_l}(x_1) \prod_{w=2}^{d+1} \sum_{t=r_{w-1}}^{r_w-1} (-1)^{r_w-1-t} \sum_{\substack{n_\tau=t-r_{w-1} \\ n_{\tau'}=r_w-1-t}} \prod_{l=1}^{t-r_{w-1}} uvF_{\tau_l}(x_w) \prod_{l=1}^{r_w-1-t} uvF_{\tau'_l}(x_{w-1}) \prod_{w=1}^d uvf_{j_{r_w}}(x_w) \tag{19}$$

where $\sum_{D \in P_{r_d, \dots, r_2, n}}$ denotes the sum over all permutations (j_1, j_2, \dots, j_n) of $(1, 2, \dots, n)$ for $j_1 < j_2 < \dots < j_{r_1-1}, j_{r_1+1} < j_{r_1+2} < \dots < j_{r_2-1}, \dots, j_{r_d+1} < j_{r_d+2} < \dots < j_n$.

Theorem 2.4. can be expressed as Theorem 2.6.

Theorem 2.6.

$$uvf_{r_1, r_2, \dots, r_d:n}(x_1, x_2, \dots, x_d) = (n - r_d)! \sum_{P_d} \prod_{l=1}^{r_1-1} uvF_{j_l}(x_1) \prod_{w=2}^{d+1} \sum_{t=r_{w-1}}^{r_w-1} (-1)^{r_w-1-t} \sum_{\substack{n_\tau=t-r_{w-1} \\ n_{\tau'}=r_w-1-t}} \prod_{l=1}^{t-r_{w-1}} uvF_{\tau_l}(x_w) \prod_{l=1}^{r_w-1-t} uvF_{\tau'_l}(x_{w-1}) \prod_{w=1}^d uvf_{j_{r_w}}(x_w) \tag{20}$$

where \sum_{P_d} denotes the sum over all permutations $\{j_1, j_2, \dots, j_d\}$ of $\{1, 2, \dots, n\}$.

Theorem 2.4. can be expressed as Theorem 2.7. using (5) and (6).

Theorem 2.7.

$$\begin{aligned}
 {}_{uv}f_{r_1, r_2, \dots, r_d; n}(x_1, x_2, \dots, x_d) &= \sum \sum n! D[{}_{uv}F^S(x_1)]^{r_1-1} \\
 &\cdot \prod_{w=2}^{d+1} \sum_{t=r_{w-1}}^{r_w-1} (-1)^{r_w-1-t} \binom{r_w-r_{w-1}-1}{t-r_{w-1}} [{}_{uv}F^S(x_w)]^{t-r_{w-1}} [{}_{uv}F^S(x_{w-1})]^{r_w-1-t} \prod_{w=1}^d {}_{uv}f^S(x_w)
 \end{aligned} \tag{21}$$

3. Results

In this section, we derive some results for the probability density functions of order statistics of innid truncated continuous random variables. We will express the following result for pdf of the r th order statistics from innid variables from truncated distribution.

Result 3.1.

$$\begin{aligned}
 {}_{uv}f_{r;n}(x) &= \frac{1}{(r-1)!(n-r)!} \sum_{t=r}^n (-1)^{n-t} \binom{n-r}{t-r} \sum_{n_\xi=n+r-t} (t-r)! \\
 &\sum_{n_\xi=n+r-1-t} \frac{\text{per}[{}_{uv}F(x)][\zeta/\cdot] \text{per}[{}_{uv}f(x)][\zeta' / \cdot]}{n+r-1-t} \\
 &= \frac{1}{(r-1)!(n-r)!} \sum_{t=r}^n (-1)^{n-t} \binom{n-r}{t-r} \sum_{n_\xi=n+r-t} (t-r)! \sum_{n_\xi=n+r-1-t} (n+r-1-t)! \prod_{l=1}^{n+r-1-t} {}_{uv}F_{\zeta^l}(x) {}_{uv}f_{\zeta^l}(x) \\
 &= \sum \sum r \binom{n}{r} \sum_{t=r}^n (-1)^{n-t} \binom{n-r}{t-r} [{}_{uv}F^S(x)]^{n+r-1-t} {}_{uv}f^S(x) \\
 &= \frac{1}{(r-1)!(n-r)!} \sum_P \prod_{l=1}^{r-1} {}_{uv}F_{j_l}(x) \sum_{t=r}^n (-1)^{n-t} \sum_{n_{\tau_l}=n-t} \prod_{l=1}^{n-t} {}_{uv}F_{\tau_l}(x) {}_{uv}f_{j_r}(x) \\
 &= (r-1)!(n-r)! \sum_D \prod_{l=1}^{r-1} {}_{uv}F_{j_l}(x) \sum_{t=r}^n (-1)^{n-t} \sum_{n_{\tau_l}=n-t} \prod_{l=1}^{n-t} {}_{uv}F_{\tau_l}(x) {}_{uv}f_{j_r}(x) \\
 &= (n-r)! \sum_P \prod_{l=1}^{r-1} {}_{uv}F_{j_l}(x) \sum_{t=r}^n (-1)^{n-t} \sum_{n_{\tau_l}=n-t} \prod_{l=1}^{n-t} {}_{uv}F_{\tau_l}(x) {}_{uv}f_{j_r}(x) \\
 &= \sum \sum r \binom{n}{r} [{}_{uv}F^S(x)]^{r-1} \sum_{t=r}^n (-1)^{n-t} \binom{n-r}{t-r} [{}_{uv}F^S(x)]^{n-t} {}_{uv}f^S(x)
 \end{aligned} \tag{22}$$

Proof. In (7), (11), (12), (14), (19), (20) and (21), if $d=1$, (22) is obtained.

We will express the following result for pdf of the minimum order statistics from innid variables from truncated distribution.

Result 3.2.

$${}_{uv}f_{1;n}(x) = \frac{1}{(n-1)!} \sum_{t=1}^n (-1)^{n-t} \binom{n-1}{t-1} \sum_{n_\xi=n+1-t} (t-1)! \sum_{n_\xi=n-t} \frac{\text{per}[{}_{uv}F(x)][\zeta/\cdot] \text{per}[{}_{uv}f(x)][\zeta' / \cdot]}{n-t}$$

$$\begin{aligned}
 &= \frac{1}{(n-1)!} \sum_{t=1}^n (-1)^{n-t} \binom{n-1}{t-1} \sum_{n_\zeta=n+1-t} (t-1)! \sum_{n_\zeta=n-t} (n-t)! \prod_{l=1}^{n-t} {}_{uv}F_{\zeta^l}(x) {}_{uv}f_{\zeta^l}(x) \\
 &= \sum_P \sum_{t=1}^n n \sum_{t=1}^n (-1)^{n-t} \binom{n-1}{t-1} [{}_{uv}F^S(x)]^{n-t} {}_{uv}f^S(x) \\
 &= \frac{1}{(n-1)!} \sum_P \sum_{t=1}^n (-1)^{n-t} \sum_{n_\zeta=n-t} \prod_{l=1}^{n-t} {}_{uv}F_{\zeta^l}(x) {}_{uv}f_{j_l}(x) \\
 &= (n-1)! \sum_{D P_1} \sum_{t=1}^n (-1)^{n-t} \sum_{n_\zeta=n-t} \prod_{l=1}^{n-t} {}_{uv}F_{\zeta^l}(x) {}_{uv}f_{j_l}(x) \\
 &= (n-1)! \sum_{P_1} \sum_{t=1}^n (-1)^{n-t} \sum_{n_\zeta=n-t} \prod_{l=1}^{n-t} {}_{uv}F_{\zeta^l}(x) {}_{uv}f_{j_l}(x)
 \end{aligned} \tag{23}$$

Proof. In (22), if $r = 1$, (23) is obtained.

We will express the following result for pdf of the maximum order statistics from innid variables from truncated distribution.

Result 3.3.

$$\begin{aligned}
 {}_{uv}f_{n:n}(x) &= \frac{1}{(n-1)!} \sum_{n_\zeta=n-1} \text{per}[{}_{uv}F(x)] \binom{n-1}{n-1} \text{per}[{}_{uv}f(x)] \binom{n-1}{1} \\
 &= \sum_{n_\zeta=n-1} \prod_{l=1}^{n-1} {}_{uv}F_{\zeta^l}(x) {}_{uv}f_{\zeta^l}(x) \\
 &= \sum_P \sum_{t=1}^n n [{}_{uv}F^S(x)]^{n-1} {}_{uv}f^S(x) \\
 &= \frac{1}{(n-1)!} \sum_P \prod_{l=1}^{n-1} {}_{uv}F_{j_l}(x) {}_{uv}f_{j_n}(x) \\
 &= (n-1)! \sum_{D P_n} \prod_{l=1}^{n-1} {}_{uv}F_{j_l}(x) {}_{uv}f_{j_n}(x) \\
 &= \sum_{P_n} \prod_{l=1}^{n-1} {}_{uv}F_{j_l}(x) {}_{uv}f_{j_n}(x)
 \end{aligned} \tag{24}$$

Proof. In (22), if $r = 1$, (24) is obtained.

We will express the following result for the joint pdf of $X_{1:n}$ and $X_{n:n}$ of the order statistics from innid variables from truncated distribution.

Result 3.4.

$$\begin{aligned}
 {}_{uv}f_{1,n:n}(x_1, x_2) &= \frac{1}{(n-2)!} \sum_{t=1}^{n-1} (-1)^{n-1-t} \binom{n-2}{t-1} \sum_{n_{\zeta_1}=n-t} \prod_{w=1}^2 \sum_{n_{\zeta_w}} \text{per}[{}_{uv}F(x_w)] \binom{n-2}{r_{w+1}-r_{w-1}-1-t_w+t_{w-1}} \text{per}[{}_{uv}f(x_w)] \binom{n-2}{1} \\
 &= \frac{1}{(n-2)!} \sum_{t=1}^{n-1} (-1)^{n-1-t} \binom{n-2}{t-1} \sum_{n_{\zeta_1}=n-t} \prod_{w=1}^2 \sum_{n_{\zeta_w}} n_{\zeta_w}! \prod_{l=1}^{n_{\zeta_w}} {}_{uv}F_{\zeta^l}(x_w) {}_{uv}f_{\zeta^l}(x_w)
 \end{aligned}$$

$$\begin{aligned}
&= \sum \sum n(n-1) \sum_{t=1}^{n-1} (-1)^{n-1-t} \binom{n-2}{t-1} [{}_{uv}F^S(x_1)]^{n-1-t} [{}_{uv}F^S(x_2)]^{t-1} {}_{uv}f^S(x_1) {}_{uv}f^S(x_2) \\
&= \frac{1}{(n-2)!} \sum_P \sum_{t=1}^{n-1} (-1)^{n-1-t} \sum_{\substack{n_\tau=t-1 \\ n_{\tau'}=n-1-t}} \prod_{l=1}^{t-1} {}_{uv}F_{\tau_l}(x_2) \prod_{l=1}^{n-1-t} {}_{uv}F_{\tau'_l}(x_1) {}_{uv}f_{j_1}(x_1) {}_{uv}f_{j_n}(x_2) \\
&= (n-2)! \sum_{d P_{n,1}} \sum_{t=1}^{n-1} (-1)^{n-1-t} \sum_{\substack{n_\tau=t-1 \\ n_{\tau'}=n-1-t}} \prod_{l=1}^{t-1} {}_{uv}F_{\tau_l}(x_2) \prod_{l=1}^{n-1-t} {}_{uv}F_{\tau'_l}(x_1) {}_{uv}f_{j_1}(x_1) {}_{uv}f_{j_n}(x_2) \\
&= \sum_{P_n} \sum_{t=1}^{n-1} (-1)^{n-1-t} \sum_{\substack{n_\tau=t-1 \\ n_{\tau'}=n-1-t}} \prod_{l=1}^{t-1} {}_{uv}F_{\tau_l}(x_2) \prod_{l=1}^{n-1-t} {}_{uv}F_{\tau'_l}(x_1) {}_{uv}f_{j_1}(x_1) {}_{uv}f_{j_n}(x_2)
\end{aligned} \tag{25}$$

Proof. In (7), (11), (12), (14), (19), (20) and (21), if $d = 2$, $r_1 = 1$ and $r_2 = n$, (25) is obtained.

References

1. Vaughan, R. J. and Venables, W. N. 'Permanent expressions for order statistics densities', *Journal of the Royal Statistical Society, Ser. B.* (34), 308-310, 1972.
2. Balasubramanian, K., Beg, M.I. & Bapat, R.B. 'On families of distributions closed under extrema', *Sankhyâ, Ser.A.* (53), 375-388, 1991.
3. Beg, M. I. 'Recurrence relations and identities from product moments of order statistics corresponding to nonidentically distributed variables', *Sankhyâ, Ser.A.* (53), 365-374, 1991.
4. Mohie, El-Din M.M., Mahmoud, M.A.W., Abu-Youssef, S.E. and Sultan, K.S. 'Order statistics from doubly truncated linear-exponential distribution and its characterizations', *Commun. Statist.-Simul. and Comput.* (26), 281-290, 1997.
5. Cao, G. and West, M. 'Computing distributions of order statistics', *Communications in Statistics Theory and Methods*, (26), 755-764, 1997.
6. Childs, A. and Balakrishnan, N. (1998). Generalized recurrence relations for moment of order statistics from non-identical Pareto and truncated Pareto random variables with applications to robustness. *Handbook of Statistics* Vol. 16, 403-438 (North-Holland Amsterdam).
7. Balakrishnan, N. and Joshi, P. C. (1982). Moments of order statistics from doubly truncated Pareto distribution. *Journal of the Indian Statistical Association* 20, 109-117.
8. Ahmad, Abd el-baset A. 'Moments of order statistics from doubly truncated continuous distributions and characterizations', *Statistics: A journal of Theoretical and Applied Statistics* ,(35), 479-494, 2001.
9. Balakrishnan, N. 'Permanents, order statistics, outliers and robustness', *Rev. Mat. Complut.*, (20), 7-107, 2007.
10. Reiss, R. D. 'Approximate distributions of order statistics', Springer, Verlag, New York Inc., USA, 1989.
11. Balakrishnan, N. and Cohen, A. C. 'Order statistics and inference', Academic Press, Inc., San Diego, 1991.
12. Galambos, J. 'The Asymptotic Theory of Extreme Order Statistics', Robert E. Krieger Publishing Co., Inc., Malabar, Florida, 1987.
13. Barakat, H. M. and Abdelkader, Y. H., 'Computing the moments of order statistics from nonidentically distributed Weibull variables, *Journal of Computational and Applied Mathematics*, (117), 85-90. 2000.